The Study of the Probability of Overfitting and the Signal-to-Noise Ratio of the KIC_U Criterion

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Abstract

This paper shows the derivation of the probability of overfitting and the signal-to-noise ratio of the \mathbf{KIC}_{U} criterion. Comparing them with **AIC**, \mathbf{AIC}_{C} , \mathbf{AIC}_{U} , **SBC**, **KIC** and **KIC**_C on AR(1), AR(2), and AR(3) models were examined. The results show that, for small to medium sample sizes, the **KIC**_U criterion had the lowest probability of overfitting and the highest signal-to-noise ratio. However, the **SBC** criterion is the best for large sample.

1. Introduction

Model selection criteria are often compared using results from simulation studies. Often, a count of time that a selection criterion identifies the correct model is a useful measure of model selection performance. However, the more variety in models, the more unreliable can become [7]. They proposed the two theoretical properties to measure a model selection performance that is the probability of overfitting and the signal to noise ratio. The aim of this paper is to derive this two theoretical properties of the **KIC**_U criterion [3] in autoregressive model. Omitting the constant $n \ln 2\pi$, the **KIC**_U criterion for AR model at order *p* is given by

$$\mathbf{KIC}_{\mathbf{U}}(p) = n \ln \hat{s}_{p}^{2} + n \ln \left(\frac{n}{n-p}\right)$$

$$+ \frac{n[(n+p)(n-p) + (n-p-2)]}{(n-p-2)(n-p)}$$
(1)

where \hat{s}_p^2 is an unbiased estimator of error variance of the process with sample of size *n*.

2. Probability of Overfitting

When choosing the best model from the candidate models for each model selection criterion, the best model is assumed to have the lowest selection criterion value. In the case of overfitting, we assume that the true model order is k and a candidate model order is p where p=k+L and L>0, L being the amount of overfitting. If the criterion value of a model of order k+L is less than that of

order k, $\mathbf{KIC}_{U}(k+L) < \mathbf{KIC}_{U}(k)$, the candidate model with an order k+L is selected instead of the true model of order k. The model with order k+L is said to be *overfitted*. Then the probability of overfitting by L for the **KIC**_U criterion is given by

$$P\{\mathbf{KIC}_{\mathrm{U}}(k+L) < \mathbf{KIC}_{\mathrm{U}}(k)\} \quad . \tag{2}$$

The criterion that gives the lower probability of overfitting is the better one. Then we have

$$\begin{split} P \big\{ \mathbf{KIC}_{\mathrm{U}}(k+L) < \mathbf{KIC}_{\mathrm{U}}(k) \big\} \\ &= P \left\{ \left[\ln \hat{s}_{k+L}^2 + \ln \left(\frac{n}{n-k-L} \right) \right. \\ &+ \frac{\left[(n+k+L)(n-k-L) + (n-k-L-2) \right]}{(n-k-L-2)(n-k-L)} \right] \\ &\left. < \left[\ln \hat{s}_k^2 + \ln \left(\frac{n}{n-k} \right) \right. \\ &\left. + \frac{\left[(n+k)(n-k) + (n-k-2) \right]}{(n-k-2)(n-k)} \right] \right\} \end{split}$$

Since $SSE_k = (n-k)\hat{s}_k^2$ and therefore $\ln SSE_k = \ln(n-k) + \ln \hat{s}_k^2$, then

$$P\{\mathbf{KIC}_{U}(k+L) < \mathbf{KIC}_{U}(k)\}$$

$$= P\left\{ \begin{bmatrix} \ln SSE_{k+L} - \ln(n-k-L) - \ln(n-k-L) \\ + \frac{[(n+k+L)(n-k-L) + (n-k-L-2)]}{(n-k-L-2)(n-k-L)} \end{bmatrix} \right\}$$

$$< \begin{bmatrix} \ln SSE_{k} - \ln(n-k) - \ln(n-k) \\ + \frac{[(n+k)(n-k) + (n-k-2)]}{(n-k-2)(n-k)} \end{bmatrix} \right\}$$

$$P\left\{ \mathbf{KIC}_{\mathrm{U}}(k+L) < \mathbf{KIC}_{\mathrm{U}}(k) \right\}$$

= $P\left\{ \ln \frac{SSE_{k}}{SSE_{k+L}} > \left[2\ln \frac{(n-k)}{(n-k-L)} + \frac{[(n+k+L)(n-k-L)+(n-k-L-2)]]}{(n-k-L-2)(n-k-L)} - \frac{[(n+k)(n-k)+(n-k-2)]}{(n-k-2)(n-k)} \right] \right\}.$
(3)

We know that

$$\begin{cases} \frac{[(n+k+L)(n-k-L) + (n-k-L-2)]}{(n-k-L-2)(n-k-L)} \\ -\frac{[(n+k)(n-k) + (n-k-2)]}{(n-k-2)(n-k)} \\ \end{cases}$$
$$= \frac{L[2(n-1)(n-k)(n-k-L) + (n-k-2)(n-k-L-2)]}{(n-k-L-2)(n-k-L)(n-k-2)(n-k)} \end{cases}$$

then, let

$$\mathbf{A} = \frac{L[2(n-1)(n-k)(n-k-L) + (n-k-2)(n-k-L-2)]}{(n-k-L-2)(n-k-L)(n-k-2)(n-k)}$$

Substituting in (3), the following is obtained;

$$P\left\{\operatorname{KIC}_{U}(k+L) < \operatorname{KIC}_{U}(k)\right\}$$

$$= P\left\{\ln\frac{SSE_{k}}{SSE_{k+L}} > 2\ln\frac{(n-k)}{(n-k-L)} + \mathbf{A}\right\}$$

$$= P\left\{\frac{SSE_{k}}{SSE_{k+L}} > \left\{\frac{(n-k)}{(n-k-L)}\right\}^{2} \exp(\mathbf{A})\right\}$$

$$= P\left\{\frac{SSE_{k}}{SSE_{k+L}} - 1 > \left\{\frac{(n-k)}{(n-k-L)}\right\}^{2} \exp(\mathbf{A}) - 1\right\}$$

$$= P\left\{\frac{SSE_{k} - SSE_{k+L}}{SSE_{k+L}} > \left\{\frac{(n-k)}{(n-k-L)}\right\}^{2} \exp(\mathbf{A}) - 1\right\}.$$

From [7], $SSE_k - SSE_{k+L} \sim \sigma_k^2 \chi_L^2$, $SSE_{k+L} \sim \sigma_k^2 \chi_{n-k-L}^2$ and $SSE_k - SSE_{k+L}$ are independent of SSE_{k+L} . Then

$$P\left\{\operatorname{KIC}_{U}(k+L) < \operatorname{KIC}_{U}(k)\right\}$$

$$= P\left\{\frac{\chi_{L}^{2}}{\chi_{n-k-L}^{2}} > \left\{\frac{(n-k)}{(n-k-L)}\right\}^{2} \exp\left(\operatorname{\mathbf{A}}\right) - 1\right\}$$

$$= P\left\{F_{L,n-k-L} > \left(\frac{n-k-L}{L}\right) \times \left(\left\{\frac{(n-k)}{(n-k-L)}\right\}^{2} \exp\left(\operatorname{\mathbf{A}}\right) - 1\right)\right\}.$$
(4)

From [7], for a fixed *k* and *L*, and where *n* approaches infinity, $\frac{\chi^2_{n-k-L}}{(n-k-L)} \rightarrow 1$, $F_{L,n-k-L} \rightarrow \frac{\chi^2_L}{L}$, and if $\lim_{n \to \infty} f_n \to f \quad \text{then} \quad \lim_{n \to \infty} P\left\{F_{L,n-k-L} > f_n\right\}$ $\to P\left\{\chi_L^2 > fL\right\}, \text{ and } \exp(x) = 1 + x + \sum_{i=2}^{\infty} \frac{x^i}{i!}. \text{ is replaced by}$ $\exp(x) = 1 + x + o(x^2) \text{ for } 0 \le x \le 1. \text{ From (4), if}$

$$f_n = \left(\frac{n-k-L}{L}\right) \times \left(\left\{\frac{(n-k)}{(n-k-L)}\right\}^2 \exp(\mathbf{A}) - 1\right)$$

then $\lim_{n\to\infty} f_n \to 4$ and

$$\lim_{n\to\infty} P\left\{F_{L,n-k-L} > f_n\right\} \to P\left\{\chi_L^2 > 4L\right\}.$$

Therefore, the asymptotic probability of overfitting of the **KIC**_U criterion by *L* is $P\{\chi_L^2 > 4L\}$. Notice that $P\{\chi_L^2 > 4L\}$ decreases as the amount of overfitting *L* increases.

3. Signal-to-Noise Ratio

The signal-to-noise ratio is a measurement which is basically a ratio of the expectation to the standard deviation of the difference in criterion values for two models. The ratio tends to assess whether the penalty term is sufficiently strong in relation to the goodness-offit term. From the true model order k and a candidate model order p where p=k+L and L>0, the true model is considered better than a candidate model if the criterion value of a model of order k is less than that of order k+L, $\mathbf{KIC}_{\mathrm{U}}(k) < \mathbf{KIC}_{\mathrm{U}}(k+L)$. For the $\mathbf{KIC}_{\mathrm{U}}$ criterion, McQuarrie and Tsai [7] defined the signal as $E[\mathbf{KIC}_{U}(k+L) - \mathbf{KIC}_{U}(k)]$, and the noise as the standard deviation of the difference, $sd[\mathbf{KIC}_{U}(k+L) - \mathbf{KIC}_{U}(k)]$. Then the signal-to-noise ratio that the true model is selected compared with a candidate model is defined as

$$\frac{E[\mathbf{KIC}_{\mathrm{U}}(k+L) - \mathbf{KIC}_{\mathrm{U}}(k)]}{sd[\mathbf{KIC}_{\mathrm{U}}(k+L) - \mathbf{KIC}_{\mathrm{U}}(k)]}$$
(5)

The criterion that gives a higher signal-to-noise ratio is the better one. Notice that when the amount of overfitting, L, increases, the signal-to-noise ratio will increase to indicate a higher overfit.

Since $E(\ln \hat{s}_p^2) \simeq \ln \sigma_k^2 - \frac{1}{n-p}$, the following is

obtained;

$$\begin{split} E(\mathbf{KIC}_{\mathbf{U}}(k+L)) &- E(\mathbf{KIC}_{\mathbf{U}}(k)) \\ & \simeq \left\{ \left[n \ln \sigma_k^2 - \frac{n}{n-k-L} + n \ln \left(\frac{n}{n-k-L} \right) \right. \\ & \left. + \frac{n[(n+k+L)(n-k-L) + (n-k-L-2)]}{(n-k-L-2)(n-k-L)} \right] \right\} \\ & \left. - \left\{ \left[n \ln \sigma_k^2 - \frac{n}{n-k} + n \ln \left(\frac{n}{n-k} \right) \right. \\ & \left. + \frac{n[(n+k)(n-k) + (n-k-2)]}{(n-k-2)(n-k)} \right] \right\} \end{split}$$

$$\begin{split} E(\mathbf{KIC}_{U}(k+L)) &= E(\mathbf{KIC}_{U}(k)) \\ &\simeq \frac{-nL}{(n-k-L)(n-k)} + n \ln \left(\frac{n-k}{n-k-L}\right) \\ &+ \frac{nL[2(n-1)(n-k)(n-k-L) + (n-k-2)(n-k-L-2)]}{(n-k-L-2)(n-k-L)(n-k-2)(n-k)} \end{split}$$

To measure the noise;

$$sd\left[\mathbf{KIC}_{\mathbf{U}}(k+L) - \mathbf{KIC}_{\mathbf{U}}(k)\right]$$
$$= sd\left[n\ln\hat{s}_{k+L}^2 - n\ln\hat{s}_k^2 + \text{constant}\right]$$
$$= sd\left[n\ln\frac{SSE_{k+L}}{SSE_k}\right].$$

McQuarrie and Tsai [7] suggested an approximation of $Var\left(n \ln \frac{SSE_{k+L}}{SSE_k}\right)$ by $\frac{2n^2L}{(n-k-L)(n-k+2)}$. Then the

noise becomes

noise
$$\simeq \frac{n\sqrt{2L}}{\sqrt{(n-k-L)(n-k+2)}}$$
. (6)

Therefore, the signal-to-noise ratio of the KIC_{U} criterion for the AR(k) model is given by

$$\frac{E(\mathbf{KIC}_{U}(k+L)) - E(\mathbf{KIC}_{U}(k))}{sd[\mathbf{KIC}_{U}(k+L) - \mathbf{KIC}_{U}(k)]} \approx \frac{\sqrt{(n-k-L)(n-k+2)}}{n\sqrt{2L}} \times \left[\frac{-nL}{(n-k-L)(n-k)} + n\ln\left(\frac{n-k}{n-k-L}\right) + \frac{nL[2(n-1)(n-k)(n-k-L) + (n-k-2)(n-k-L-2)]}{(n-k-L-2)(n-k-L)(n-k-2)(n-k)}\right]$$
(7)

For a fixed k and L,

$$\lim_{n \to \infty} \frac{E\left[\mathbf{KIC}_{\mathrm{U}}(k+L) - \mathbf{KIC}_{\mathrm{U}}(k)\right]}{sd\left[\mathbf{KIC}_{\mathrm{U}}(k+L) - \mathbf{KIC}_{\mathrm{U}}(k)\right]} = \frac{3L}{\sqrt{2L}}$$

Therefore, the asymptotic signal-to-noise ratio of the **KIC**_U criterion for the AR model is $\frac{3L}{\sqrt{2L}}$. Notice that $\frac{3L}{\sqrt{2L}}$ increases as the amount of overfitting *L* increases.

4. Simulation Study

In this section, the probability of overfitting and the signal-to-noise ratio of the \mathbf{KIC}_{U} criterion were tested by comparing them with a number of considered criteria using simulation. The considered criteria are **AIC** [1], **AIC**_C [5], **AIC**_U [6], **SBC** [8], **KIC** [4] and **KIC**_C [2] which their probability of overfitting and signal-to-noise ratio are given in McQuarrie and Tsai [7].

AR(1), AR(2) and AR(3) models with sample sizes 25, 40, 60 and 100, ranging from small to large, were used. All results are shown in tables 1-6.

The data shown in table 1 are the probability of overfitting for the AR(1) model by each criteria for different amounts of overfitting L and different sample sizes, e.g. for KIC_U where n = 25 and L = 1 the probability of overfitting was 0.0382. This means that this criterion would select the model whose order is higher by one order than true model, AR(1), with a probability of 0.0382. It can be seen that, where n = 25, $\mathbf{KIC}_{\mathbf{U}}$ is the best among the selection criteria as it had the lowest probability of overfitting; 0.0382 and 0.0128 for L = 1 and 2 respectively. Where n = 40, **KIC**_U is the best with the lowest probability of overfitting; 0.0411, 0.0150 and 0.0053 for L = 1, 2 and 3 respectively. Where n = 60, **KIC**_U performs the best with the lowest probability of overfitting; 0.0427, 0.0161 and 0.0060 for L = 1, 2 and 3 respectively. Where n = 100, the best criterion is **SBC** with the lowest probability of overfitting; 0.0341, 0.0115 and 0.0040 for L = 1, 2 and 3 respectively.

The data shown in table 4 are the signal-to-noise ratios for the AR(1) model for each criteria using different amounts of overfitting L for different sample sizes, e.g. for KIC_U , where n = 25 and L = 1, the signalto-noise ratio was 2.5325. This means that this criterion will select the model whose order is higher by one order than the true model, AR(1), with a signal-to-noise ratio of 2.5325. It was found that, where n = 25, KIC_U has the highest signal-to-noise ratio; 2.5325 and 3.6494 for L = 1and 2 respectively. Where n = 40, the highest signal-tonoise ratio was generated by KIC_U; 2.3594, 3.3711 and 4.1729 for L = 1, 2 and 3 respectively. Where n = 60, **KIC**_U generated the highest signal-to-noise ratio; 2.2738, 3.2362 and 3.9894 for L = 1, 2 and 3 respectively. Where n = 100, the highest signal-to-noise ratio was generated by **SBC**; 2.5182, 3.5378 and 4.3041 for L = 1, 2 and 3 respectively.

 Table 1 The probability of overfitting for AR(1) model

п	T		р	robabili	ity of overfitting			
	L	AIC	AICc	AICu	SBC	KIC	KICc	KICu
25	1	0.1796	0.1262	0.0887	0.0887	0.1003	0.0677	0.0382
	2	0.1720	0.0907	0.0589	0.0589	0.0714	0.0334	0.0128
23	3	0.1608	0.0602	0.0403	0.0403	0.0519	0.0155	0.0041
	4	0.1524	0.0380	0.0291	0.0291	0.0396	0.0069	0.0013
	1	0.1709	0.1383	0.0629	0.0629	0.0935	0.0738	0.0411
40	2	0.1572	0.1077	0.0330	0.0330	0.0623	0.0396	0.0150
40	3	0.1402	0.0790	0.0178	0.0178	0.0419	0.0205	0.0053
	4	0.1260	0.0565	0.0101	0.0101	0.0290	0.0104	0.0019
	1	0.1662	0.1448	0.0476	0.0476	0.0900	0.0771	0.0427
60	2	0.1496	0.1170	0.0205	0.0205	0.0578	0.0431	0.0161
00	3	0.1300	0.0898	0.0090	0.0090	0.0372	0.0234	0.0060
	4	0.1133	0.0678	0.0042	0.0042	0.0245	0.0126	0.0022
	1	0.1626	0.1499	0.0341	0.0341	0.0872	0.0796	0.0438
100	2	0.1437	0.1244	0.0115	0.0115	0.0545	0.0458	0.0170
100	3	0.1223	0.0985	0.0040	0.0040	0.0338	0.0257	0.0066
	4	0.1041	0.0771	0.0014	0.0014	0.0213	0.0144	0.0025

Table 3 The probability of overfitting for AR(3) model

10	L	probability of overfitting						
n		AIC	AIC _C	AIC _U	SBC	KIC	KICc	KICu
25	1	0.2002	0.1075	0.1041	0.1041	0.1167	0.0581	0.0330
	2	0.2019	0.0695	0.0761	0.0761	0.0907	0.0256	0.0099
23	3	0.1972	0.0412	0.0570	0.0570	0.0716	0.0105	0.0028
	4	0.1945	0.0229	0.0449	0.0449	0.0589	0.0041	0.0007
	1	0.1827	0.1271	0.0704	0.0704	0.1027	0.0682	0.0381
40	2	0.1738	0.0941	0.0396	0.0396	0.0724	0.0346	0.0131
40	3	0.1595	0.0656	0.0229	0.0229	0.0513	0.0169	0.0044
	4	0.1472	0.0446	0.0138	0.0138	0.0374	0.0081	0.0015
	1	0.1738	0.1375	0.0516	0.0516	0.0957	0.0734	0.0407
60	2	0.1599	0.1079	0.0234	0.0234	0.0639	0.0397	0.0149
00	3	0.1417	0.0805	0.0109	0.0109	0.0426	0.0209	0.0054
	4	0.1258	0.0591	0.0053	0.0053	0.0290	0.0109	0.0019
	1	0.1670	0.1456	0.0360	0.0360	0.0905	0.0775	0.0427
100	2	0.1496	0.1190	0.0126	0.0126	0.0578	0.0438	0.0163
100	3	0.1288	0.0928	0.0045	0.0045	0.0367	0.0242	0.0062
	4	0.1109	0.0716	0.0017	0.0017	0.0236	0.0133	0.0023

 Table 2 The probability of overfitting for AR(2) model

n	L	probability of overfitting						
		AIC	AICc	AICu	SBC	KIC	KICc	KICu
	1	0.1896	0.1169	0.0961	0.0961	0.1081	0.0630	0.0356
25	2	0.1864	0.0800	0.0669	0.0669	0.0805	0.0295	0.0113
23	3	0.1781	0.0503	0.0479	0.0479	0.0610	0.0129	0.0034
	4	0.1723	0.0299	0.0362	0.0362	0.0484	0.0054	0.0010
	1	0.1767	0.1328	0.0665	0.0665	0.0980	0.0710	0.0396
40	2	0.1653	0.1009	0.0362	0.0362	0.0672	0.0371	0.0140
40	3	0.1496	0.0722	0.0202	0.0202	0.0463	0.0187	0.0048
	4	0.1362	0.0504	0.0118	0.0118	0.0330	0.0092	0.0017
	1	0.1699	0.1412	0.0496	0.0496	0.0928	0.0753	0.0417
60	2	0.1546	0.1125	0.0219	0.0219	0.0608	0.0414	0.0155
00	3	0.1357	0.0851	0.0099	0.0099	0.0398	0.0221	0.0057
	4	0.1194	0.0634	0.0047	0.0047	0.0266	0.0117	0.0021
	1	0.1648	0.1478	0.0350	0.0350	0.0889	0.0785	0.0433
100	2	0.1466	0.1217	0.0120	0.0120	0.0561	0.0448	0.0166
100	3	0.1255	0.0956	0.0042	0.0042	0.0352	0.0249	0.0064
	4	0.1074	0.0743	0.0015	0.0015	0.0224	0.0139	0.0024

 Table 4 Signal to noise ratio for AR(1) model

п	T							
	L	AIC	AIC _C	AIC	SBC	KIC	KIC _C	KIC _U
	1	0.6161	1.0293	1.7652	1.4591	1.3077	1.7616	2.5325
25	2	0.8275	1.5233	2.5638	1.9936	1.7842	2.5067	3.6496
25	3	0.9589	1.9557	3.2295	2.3541	2.1036	3.0909	4.5601
1	4	1.0428	2.3719	3.8418	2.6151	2.3328	3.5957	5.3789
	1	0.6517	0.8906	1.6156	1.8301	1.3495	1.6143	2.3594
40	2	0.8952	1.2938	2.3190	2.5397	1.8690	2.2880	3.3711
40	3	1.0637	1.6285	2.8839	3.0504	2.2401	2.8087	4.1729
1	4	1.1901	1.9338	3.3831	3.4520	2.5294	3.2511	4.8716
	1	0.6707	0.8236	1.5425	2.1388	1.3717	1.5420	2.2738
(0)	2	0.9312	1.1852	2.2019	2.9895	1.9140	2.1828	3.2362
00	3	1.1192	1.4774	2.7225	3.6178	2.3123	2.6759	3.9894
1	4	1.2676	1.7366	3.1743	4.1268	2.6328	3.0929	4.6373
	1	0.6855	0.7744	1.4886	2.5182	1.3890	1.4884	2.2100
100	2	0.9593	1.1064	2.1164	3.5378	1.9491	2.1056	3.1368
100	3	1.1623	1.3691	2.6061	4.3041	2.3683	2.5796	3.8558
	4	1.3275	1.5975	3.0258	4.9364	2.7128	2.9797	4.4688

Table 5 Signal to noise ratio for AR(2) model

п	T	signal to noise ratio						
	L	AIC	AIC _C	AIC _U	SBC	KIC	KICc	KIC _U
	1	0.5567	1.1253	1.8624	1.3652	1.2201	1.8569	2.5792
25	2	0.7434	1.6668	2.7090	1.8605	1.6599	2.6439	3.6435
23	3	0.8555	2.1422	3.4181	2.1907	1.9510	3.2623	4.4591
	4	0.9231	2.6014	4.0736	2.4258	2.1559	3.7982	5.1474
	1	0.6153	0.9392	1.6646	1.7639	1.2954	1.6628	2.3791
40	2	0.8438	1.3646	2.3905	2.4460	1.7925	2.3570	3.3569
40	3	1.0006	1.7181	2.9743	2.9355	2.1463	2.8937	4.1019
	4	1.1171	2.0407	3.4909	3.3191	2.4210	3.3500	4.7258
	1	0.6467	0.8526	1.5718	2.0901	1.3359	1.5710	2.2842
60	2	0.8973	1.2270	2.2440	2.9206	1.8634	2.2239	3.2238
00	3	1.0776	1.5296	2.7751	3.5335	2.2502	2.7264	3.9404
	4	1.2195	1.7982	3.2363	4.0294	2.5611	3.1513	4.5407
	1	0.6712	0.7904	1.5046	2.4855	1.3676	1.5044	2.2151
100	2	0.9390	1.1293	2.1394	3.4916	1.9188	2.1282	3.1282
100	3	1.1375	1.3974	2.6346	4.2474	2.3312	2.6074	3.8258
	4	1.2989	1.6305	3.0590	4.8709	2.6700	3.0118	4.4114

Table 6 Signal to noise ratio for AR(3) model

п	L	signal to noise ratio						
		AIC	AICc	AIC	SBC	KIC	KICc	KICU
	1	0.4971	1.2324	1.9708	1.2711	1.1321	1.9630	2.6859
25	2	0.6588	1.8273	2.8714	1.7270	1.5352	2.7968	3.7964
23	3	0.7516	2.3516	3.6297	2.0268	1.7978	3.4538	4.6493
	4	0.8027	2.8602	4.3349	2.2358	1.9784	4.0249	5.3711
	1	0.5789	0.9908	1.7168	1.6976	1.2413	1.7143	2.4308
40	2	0.7922	1.4400	2.4665	2.3522	1.7159	2.4302	3.4301
40	3	0.9374	1.8135	3.0705	2.8204	2.0524	2.9841	4.1918
	4	1.0440	2.1546	3.6057	3.1862	2.3124	3.4552	4.8299
	1	0.6227	0.8827	1.6021	2.0414	1.3001	1.6011	2.3144
60	2	0.8633	1.2704	2.2878	2.8517	1.8127	2.2665	3.2665
00	3	1.0360	1.5839	2.8298	3.4490	2.1882	2.7788	3.9926
	4	1.1714	1.8621	3.3006	3.9318	2.4894	3.2121	4.6009
	1	0.6569	0.8067	1.5210	2.4528	1.3463	1.5207	2.2315
100	2	0.9188	1.1526	2.1629	3.4453	1.8886	2.1513	3.1513
100	3	1.1128	1.4264	2.6636	4.1907	2.2942	2.6357	3.8541
	4	1.2703	1.6643	3.0929	4.8055	2.6273	3.0445	4.4440

5. Conclusions

For small to medium sample sizes, the \mathbf{KIC}_{U} criterion had the lowest probability of overfitting and the highest signal-to-noise-ratio. However, as the sample size increases, **SBC** had the lowest probability of overfitting and the highest signal-to-noise-ratio.

Therefore, for small to medium sample sizes, the KIC_U performed the best. However, the SBC criterion is the best for large sample.

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